

Lecture 30

In this lecture we continue our study of conditional probability, looking at Bayes' Theorem.

Recall that a partition of a set S is a way of dividing up S into disjoint nonempty subsets.

More formally, a partition of a set S is a collection of subsets $\{A_1, A_2, \dots, A_n\}^\dagger$

of S such that

$$S = A_1 \cup A_2 \cup \dots \cup A_n$$

and the A_i are nonempty and pairwise disjoint.

Now let's suppose that S is a sample space. Let B be any event on S .

Then we have the following theorem.

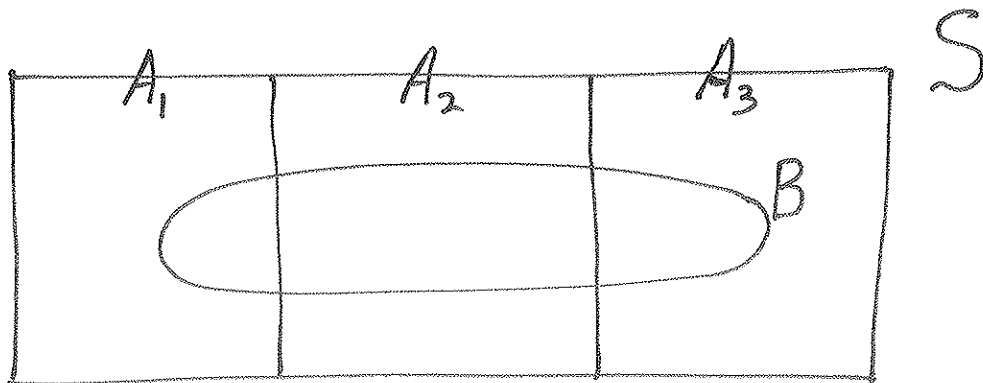
[†] In some situations n can be infinite.

Theorem

Let $\{A_1, A_2, \dots, A_n\}$ be a partition of a sample space S . Then

$$\begin{aligned}\Pr(B) &= \Pr(A_1 \cap B) + \Pr(A_2 \cap B) + \dots + \Pr(A_n \cap B) \\ &= \sum_{j=1}^n \Pr(A_j \cap B).\end{aligned}$$

Here's an illustration of the situation when $n = 3$.



Now recall that conditional probabilities are defined using intersections:

$$\Pr(D|C) = \frac{\Pr(C \cap D)}{\Pr(C)}$$

Rearranging this gives:

$$\Pr(C \cap D) = \Pr(D|C) \cdot \Pr(C)$$

We can apply this to the intersections given in the previous theorem:

$$\Pr(A_j \cap B) = \Pr(B|A_j) \cdot \Pr(A_j)$$

So the theorem can be restated as follows:

$$\begin{aligned}\Pr(B) &= \Pr(B|A_1) \cdot \Pr(A_1) + \dots + \Pr(B|A_n) \cdot \Pr(A_n) \\ &= \sum_{j=1}^n \Pr(B|A_j) \cdot \Pr(A_j)\end{aligned}$$

Then for each $i = 1, 2, \dots, n$ we have that

$$\Pr(A_i|B) = \frac{\Pr(A_i \cap B)}{\Pr(B)}.$$

Putting everything together gives us the following theorem.

Bayes' Theorem

If $\{A_1, A_2, \dots, A_n\}$ is a partition of a sample space S , and B is any event, then for each $i = 1, 2, \dots, n$ we have that

$$\Pr(A_i | B) = \frac{\Pr(B|A_i) \cdot \Pr(A_i)}{\sum_{j=1}^n \Pr(B|A_j) \cdot \Pr(A_j)}$$

E.g.

A die is cast.

$$S = \{1, 2, 3, 4, 5, 6\}$$

$$A_1 = \{1, 2\}, A_2 = \{3, 4, 5\}, A_3 = \{6\}$$

Let B be the event that "an even number is scored".

(i) Given that B occurs, what's the probability that the score is in A_1 ?

(ii) Repeat for A_2 .

(iii) Repeat for A_3 .

Solⁿ

Firstly we calculate the probability that each A_i occurs when a die is cast:

$$\Pr(A_1) = \frac{2}{6} = \frac{1}{3}$$

$$\Pr(A_2) = \frac{3}{6} = \frac{1}{2}$$

$$\Pr(A_3) = \frac{1}{6}$$

Now we calculate the conditional probability of B given each A_i :

$$\Pr(B|A_1) = \frac{\Pr(A_1 \cap B)}{\Pr(A_1)} = \frac{1}{2}$$

$$\Pr(B|A_2) = \frac{\Pr(A_2 \cap B)}{\Pr(A_2)} = \frac{1}{3}$$

$$\Pr(B|A_3) = \frac{\Pr(A_3 \cap B)}{\Pr(A_3)} = \frac{1}{1} = 1$$

So we can answer each part of the question:

(i)

$$\begin{aligned} \Pr(A_1|B) &= \frac{\Pr(B|A_1) \cdot \Pr(A_1)}{\Pr(B|A_1) \cdot \Pr(A_1) + \Pr(B|A_2) \cdot \Pr(A_2) + \Pr(B|A_3) \cdot \Pr(A_3)} \\ &= \frac{(\frac{1}{2})(\frac{1}{3})}{(\frac{1}{2})(\frac{1}{3}) + (\frac{1}{3})(\frac{1}{2}) + (1)(\frac{1}{6})} \\ &= \frac{(\frac{1}{6})}{(\frac{1}{6}) + (\frac{1}{6}) + (\frac{1}{6})} \\ &= \frac{(\frac{1}{6})}{(\frac{3}{6})} = \frac{1}{3} \end{aligned}$$

(ii)

$$\begin{aligned} \Pr(A_2|B) &= \frac{\Pr(B|A_2) \cdot \Pr(A_2)}{\Pr(B|A_1) \cdot \Pr(A_1) + \Pr(B|A_2) \cdot \Pr(A_2) + \Pr(B|A_3) \cdot \Pr(A_3)} \\ &= \frac{(\frac{1}{3})(\frac{1}{2})}{(\frac{3}{6})} = \frac{(\frac{1}{6})}{(\frac{3}{6})} = \frac{1}{3} \end{aligned}$$

(iii) $\Pr(A_3|B) = ?$

The new numerator is $\Pr(B|A_3) \cdot \Pr(A_3)$ which equals $(1) \cdot (\frac{1}{6}) = \frac{1}{6}$. The denominator is as in (i) and (ii). So the

result is that

$$\Pr(A_3|B) = \frac{\binom{1}{6}}{\binom{3}{6}} = \frac{1}{3}.$$

Notice that the denominator is the same in every case.

Let's have another look at Bayes' Theorem. The numerator is the probability of $A_i \cap B$, and the denominator is just the probability of B .

$$\Pr(A_i|B) = \frac{\Pr(B|A_i) \cdot \Pr(A_i)}{\sum_{j=1}^n \Pr(B|A_j) \cdot \Pr(A_j)}$$

$\Pr(A_i \cap B)$

$\Pr(B)$

Notice that what Bayes' Theorem enables us to do is to reverse the order in conditional probabilities.

We use $\Pr(B|A_1)$, $\Pr(B|A_2)$, ..., $\Pr(B|A_n)$ to help us evaluate $\Pr(A_1|B)$, $\Pr(A_2|B)$ and so on.

E.g. A test for a disease correctly detects that a person has the disease with probability .9, and correctly detects that a person doesn't have the disease with probability .9. If only 1% of the population has the disease, what is the probability that a person shown by the test as having the disease actually does have it?

Solⁿ

Let Y_d mean "yes, has the disease",

N_d mean "no, doesn't have the disease",

Y_t mean "yes, tests positive" and

N_t mean "no, tests negative".

Then

$$\Pr(Y_d) = .01 \text{ and } \Pr(N_d) = .99.$$

$$\text{Also, } \Pr(Y_t | Y_d) = .9 = \Pr(N_t | N_d).$$

So

$$\Pr(Y_d | Y_t) = \frac{\Pr(Y_t | Y_d) \cdot \Pr(Y_d)}{\Pr(Y_t | Y_d) \cdot \Pr(Y_d) + \Pr(Y_t | N_d) \cdot \Pr(N_d)}$$

$$\text{and since } \Pr(Y_t | N_d) = 1 - \Pr(N_t | N_d) = .1$$

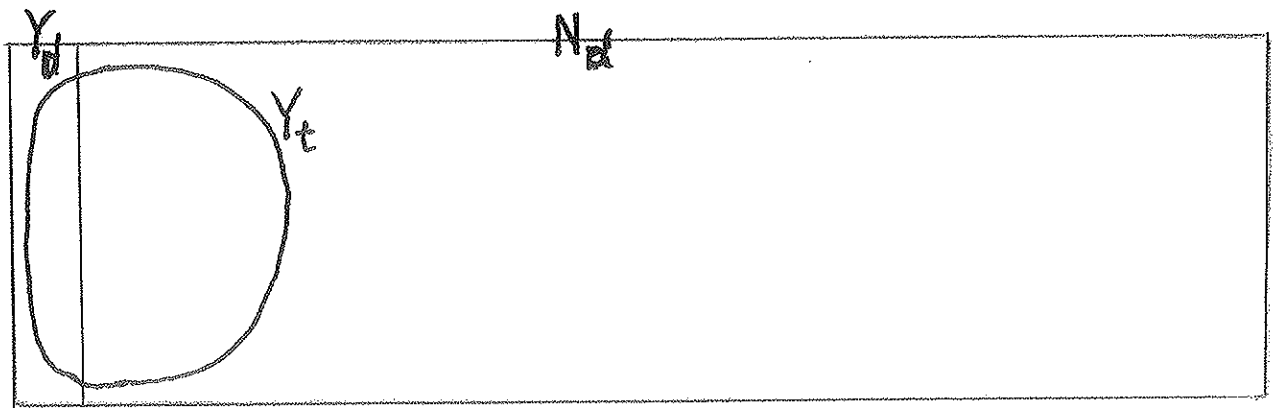
we have

$$\begin{aligned} \Pr(Y_d | Y_t) &= \frac{(.9)(.01)}{(.9)(.01) + (.1)(.99)} \\ &= \frac{.009}{.009 + .099} \\ &= \frac{.009}{.108} = \frac{9}{108} = \frac{1}{12} \end{aligned}$$

So only $\frac{1}{12}$ of those tested positive actually has the disease.

How do we explain this result?

The diagram below may help. Although test outcome Y_t occupies 90% of Y_d and only 10% of N_d , because Y_d (the sub-population of people with the disease) is so small it only occupies $\frac{1}{12}$ of Y_t .



(Not quite to scale)

So Bayes' Theorem enables us to reach a conclusion which, although possibly surprising, is correct.